

NAG Library Routine Document

G02QGF

Note: before using this routine, please read the Users' Note for your implementation to check the interpretation of ***bold italicised*** terms and other implementation-dependent details.

Note: *this routine uses optional parameters to define choices in the problem specification and in the details of the algorithm. If you wish to use default settings for all of the optional parameters, you need only read Sections 1 to 10 of this document. If, however, you wish to reset some or all of the settings please refer to Section 11 for a detailed description of the algorithm, to Section 12 for a detailed description of the specification of the optional parameters and to Section 13 for a detailed description of the monitoring information produced by the routine.*

1 Purpose

G02QGF performs a multiple linear quantile regression. Parameter estimates and, if required, confidence limits, covariance matrices and residuals are calculated. G02QGF may be used to perform a weighted quantile regression. A simplified interface for G02QGF is provided by G02QFF.

2 Specification

```
SUBROUTINE G02QGF (SORDER, INTCPT, WEIGHT, N, M, DAT, LDDAT, ISX, IP, Y,      &
                   WT, NTAU, TAU, DF, B, BL, BU, CH, RES, IOPTS, OPTS,      &
                   STATE, INFO, IFAIL)

INTEGER              SORDER, N, M, LDDAT, ISX(M), IP, NTAU, IOPTS(*),      &
                   STATE(*), INFO(NTAU), IFAIL
REAL (KIND=nag_wp)  DAT(LDDAT,*), Y(N), WT(*), TAU(NTAU), DF,              &
                   B(IP,NTAU), BL(IP,*), BU(IP,*), CH(IP,IP,*),          &
                   RES(N,*), OPTS(*)
CHARACTER(1)         INTCPT, WEIGHT
```

3 Description

Given a vector of n observed values, $y = \{y_i : i = 1, 2, \dots, n\}$, an $n \times p$ design matrix X , a column vector, x , of length p holding the i th row of X and a quantile $\tau \in (0, 1)$, G02QGF estimates the p -element vector β as the solution to

$$\underset{\beta \in \mathbb{R}^p}{\text{minimize}} \sum_{i=1}^n \rho_{\tau}(y_i - x_i^T \beta) \quad (1)$$

where ρ_{τ} is the piecewise linear loss function $\rho_{\tau}(z) = z(\tau - I(z < 0))$, and $I(z < 0)$ is an indicator function taking the value 1 if $z < 0$ and 0 otherwise. Weights can be incorporated by replacing X and y with WX and Wy respectively, where W is an $n \times n$ diagonal matrix. Observations with zero weights can either be included or excluded from the analysis; this is in contrast to least squares regression where such observations do not contribute to the objective function and are therefore always dropped.

G02QGF uses the interior point algorithm of Portnoy and Koenker (1997), described briefly in Section 11, to obtain the parameter estimates $\hat{\beta}$, for a given value of τ .

Under the assumption of Normally distributed errors, Koenker (2005) shows that the limiting covariance matrix of $\hat{\beta} - \beta$ has the form

$$\Sigma = \frac{\tau(1-\tau)}{n} H_n^{-1} J_n H_n^{-1}$$

where $J_n = n^{-1} \sum_{i=1}^n x_i x_i^T$ and H_n is a function of τ , as described below. Given an estimate of the covariance matrix, $\hat{\Sigma}$, lower ($\hat{\beta}_L$) and upper ($\hat{\beta}_U$) limits for an $(100 \times \alpha)\%$ confidence interval can be

calculated for each of the p parameters, via

$$\hat{\beta}_{Li} = \hat{\beta}_i - t_{n-p,(1+\alpha)/2} \sqrt{\hat{\Sigma}_{ii}}, \hat{\beta}_{Ui} = \hat{\beta}_i + t_{n-p,(1+\alpha)/2} \sqrt{\hat{\Sigma}_{ii}}$$

where $t_{n-p,0.975}$ is the 97.5 percentile of the Student's t distribution with $n - k$ degrees of freedom, where k is the rank of the cross-product matrix $X^T X$.

Four methods for estimating the covariance matrix, Σ , are available:

(i) Independent, identically distributed (IID) errors

Under an assumption of IID errors the asymptotic relationship for Σ simplifies to

$$\Sigma = \frac{\tau(1-\tau)}{n} (s(\tau))^2 (X^T X)^{-1}$$

where s is the sparsity function. G02QGF estimates $s(\tau)$ from the residuals, $r_i = y_i - x_i^T \hat{\beta}$ and a bandwidth h_n .

(ii) Powell Sandwich

Powell (1991) suggested estimating the matrix H_n by a kernel estimator of the form

$$\hat{H}_n = (nc_n)^{-1} \sum_{i=1}^n K\left(\frac{r_i}{c_n}\right) x_i x_i^T$$

where K is a kernel function and c_n satisfies $\lim_{n \rightarrow \infty} c_n \rightarrow 0$ and $\lim_{n \rightarrow \infty} \sqrt{n}c_n \rightarrow \infty$. When the Powell method is chosen, G02QGF uses a Gaussian kernel (i.e., $K = \phi$) and sets

$$c_n = \min(\sigma_r, (q_{r3} - q_{r1})/1.34) \times (\Phi^{-1}(\tau + h_n) - \Phi^{-1}(\tau - h_n))$$

where h_n is a bandwidth, σ_r , q_{r1} and q_{r3} are, respectively, the standard deviation and the 25% and 75% quantiles for the residuals, r_i .

(iii) Hendricks–Koenker Sandwich

Koenker (2005) suggested estimating the matrix H_n using

$$\hat{H}_n = n^{-1} \sum_{i=1}^n \left[\frac{2h_n}{x_i^T (\hat{\beta}(\tau + h_n) - \hat{\beta}(\tau - h_n))} \right] x_i x_i^T$$

where h_n is a bandwidth and $\hat{\beta}(\tau + h_n)$ denotes the parameter estimates obtained from a quantile regression using the $(\tau + h_n)$ th quantile. Similarly with $\hat{\beta}(\tau - h_n)$.

(iv) Bootstrap

The last method uses bootstrapping to either estimate a covariance matrix or obtain confidence intervals for the parameter estimates directly. This method therefore does not assume Normally distributed errors. Samples of size n are taken from the paired data $\{y_i, x_i\}$ (i.e., the independent and dependent variables are sampled together). A quantile regression is then fitted to each sample resulting in a series of bootstrap estimates for the model parameters, β . A covariance matrix can then be calculated directly from this series of values. Alternatively, confidence limits, $\hat{\beta}_L$ and $\hat{\beta}_U$, can be obtained directly from the $(1 - \alpha)/2$ and $(1 + \alpha)/2$ sample quantiles of the bootstrap estimates.

Further details of the algorithms used to calculate the covariance matrices can be found in Section 11.

All three asymptotic estimates of the covariance matrix require a bandwidth, h_n . Two alternative methods for determining this are provided:

(i) Sheather–Hall

$$h_n = \left(\frac{1.5(\Phi^{-1}(\alpha_b)\phi(\Phi^{-1}(\tau)))^2}{n(2\Phi^{-1}(\tau) + 1)} \right)^{\frac{1}{5}}$$

for a user-supplied value α_b ,

(ii) Bofinger

$$h_n = \left(\frac{4.5(\phi(\Phi^{-1}(\tau)))^4}{n(2\Phi^{-1}(\tau) + 1)^2} \right)^{\frac{1}{5}}$$

G02QGF allows optional arguments to be supplied via the IOPTS and OPTS arrays (see Section 12 for details of the available options). Prior to calling G02QGF the optional parameter arrays, IOPTS and OPTS must be initialized by calling G02ZKF with OPTSTR set to **Initialize** = G02QGF (see Section 12 for details on the available options). If bootstrap confidence limits are required (**Interval Method** = BOOTSTRAP XY) then one of the random number initialization routines G05KFF (for a repeatable analysis) or G05KGF (for an unrepeatable analysis) must also have been previously called.

4 References

- Koenker R (2005) *Quantile Regression* Econometric Society Monographs, Cambridge University Press, New York
- Mehrotra S (1992) On the implementation of a primal-dual interior point method *SIAM J. Optim.* **2** 575–601
- Nocedal J and Wright S J (1999) *Numerical Optimization* Springer Series in Operations Research, Springer, New York
- Portnoy S and Koenker R (1997) The Gaussian hare and the Laplacian tortoise: computability of squared-error versus absolute error estimators *Statistical Science* **4** 279–300
- Powell J L (1991) Estimation of monotonic regression models under quantile restrictions *Nonparametric and Semiparametric Methods in Econometrics* Cambridge University Press, Cambridge

5 Arguments

- 1: SORDER – INTEGER *Input*
On entry: determines the storage order of variates supplied in DAT.
Constraint: SORDER = 1 or 2.
- 2: INTCPT – CHARACTER(1) *Input*
On entry: indicates whether an intercept will be included in the model. The intercept is included by adding a column of ones as the first column in the design matrix, X .
 INTCPT = 'Y'
 An intercept will be included in the model.
 INTCPT = 'N'
 An intercept will not be included in the model.
Constraint: INTCPT = 'N' or 'Y'.
- 3: WEIGHT – CHARACTER(1) *Input*
On entry: indicates if weights are to be used.
 WEIGHT = 'W'
 A weighted regression model is fitted to the data using weights supplied in array WT.

WEIGHT = 'U'

An unweighted regression model is fitted to the data and array WT is not referenced.

Constraint: WEIGHT = 'U' or 'W'.

4: N – INTEGER *Input*

On entry: the total number of observations in the dataset. If no weights are supplied, or no zero weights are supplied or observations with zero weights are included in the model then $N = n$. Otherwise $N = n +$ the number of observations with zero weights.

Constraint: $N \geq 2$.

5: M – INTEGER *Input*

On entry: m , the total number of variates in the dataset.

Constraint: $M \geq 0$.

6: DAT(LDDAT,*) – REAL (KIND=nag_wp) array *Input*

Note: the second dimension of the array DAT must be at least M if SORDER = 1 and at least N if SORDER = 2.

On entry: the i th value for the j th variate, for $i = 1, 2, \dots, N$ and $j = 1, 2, \dots, M$, must be supplied in

DAT(i, j) if SORDER = 1, and

DAT(j, i) if SORDER = 2.

The design matrix X is constructed from DAT, ISX and INTCPT.

7: LDDAT – INTEGER *Input*

On entry: the first dimension of the array DAT as declared in the (sub)program from which G02QGF is called.

Constraints:

if SORDER = 1, $LDDAT \geq N$;

otherwise $LDDAT \geq M$.

8: ISX(M) – INTEGER array *Input*

On entry: indicates which independent variables are to be included in the model.

$ISX(j) = 0$

The j th variate, supplied in DAT, is not included in the regression model.

$ISX(j) = 1$

The j th variate, supplied in DAT, is included in the regression model.

Constraints:

$ISX(j) = 0$ or 1 , for $j = 1, 2, \dots, M$;

if INTCPT = 'Y', exactly $IP - 1$ values of ISX must be set to 1;

if INTCPT = 'N', exactly IP values of ISX must be set to 1.

9: IP – INTEGER *Input*

On entry: p , the number of independent variables in the model, including the intercept, see INTCPT, if present.

Constraints:

$1 \leq IP < N$;

if INTCPT = 'Y', $1 \leq IP \leq M + 1$;

if INTCPT = 'N', $1 \leq IP \leq M$.

- 10: Y(N) – REAL (KIND=nag_wp) array Input
On entry: y , the observations on the dependent variable.
- 11: WT(*) – REAL (KIND=nag_wp) array Input
Note: the dimension of the array WT must be at least N if WEIGHT = 'W'.
On entry: if WEIGHT = 'W', WT must contain the diagonal elements of the weight matrix W . Otherwise WT is not referenced.
 When
Drop Zero Weights = YES
 If $WT(i) = 0.0$, the i th observation is not included in the model, in which case the effective number of observations, n , is the number of observations with nonzero weights.
 If **Return Residuals** = YES, the values of RES will be set to zero for observations with zero weights.
Drop Zero Weights = NO
 All observations are included in the model and the effective number of observations is N, i.e., $n = N$.
Constraints:
 If WEIGHT = 'W', $WT(i) \geq 0.0$, for $i = 1, 2, \dots, N$;
 The effective number of observations ≥ 2 .
- 12: NTAU – INTEGER Input
On entry: the number of quantiles of interest.
Constraint: $NTAU \geq 1$.
- 13: TAU(NTAU) – REAL (KIND=nag_wp) array Input
On entry: the vector of quantiles of interest. A separate model is fitted to each quantile.
Constraint: $\sqrt{\epsilon} < TAU(j) < 1 - \sqrt{\epsilon}$ where ϵ is the **machine precision** returned by X02AJF, for $j = 1, 2, \dots, NTAU$.
- 14: DF – REAL (KIND=nag_wp) Output
On exit: the degrees of freedom given by $n - k$, where n is the effective number of observations and k is the rank of the cross-product matrix $X^T X$.
- 15: B(IP,NTAU) – REAL (KIND=nag_wp) array Input/Output
On entry: if **Calculate Initial Values** = NO, $B(i, l)$ must hold an initial estimates for $\hat{\beta}_i$, for $i = 1, 2, \dots, IP$ and $l = 1, 2, \dots, NTAU$. If **Calculate Initial Values** = YES, B need not be set.
On exit: $B(i, l)$, for $i = 1, 2, \dots, IP$, contains the estimates of the parameters of the regression model, $\hat{\beta}$, estimated for $\tau = TAU(l)$.
 If INTCPT = 'Y', $B(1, l)$ will contain the estimate corresponding to the intercept and $B(i + 1, l)$ will contain the coefficient of the j th variate contained in DAT, where $ISX(j)$ is the i th nonzero value in the array ISX.
 If INTCPT = 'N', $B(i, l)$ will contain the coefficient of the j th variate contained in DAT, where $ISX(j)$ is the i th nonzero value in the array ISX.

- 16: BL(IP,*) – REAL (KIND=nag_wp) array Output
- Note:** the second dimension of the array BL must be at least NTAU if **Interval Method** \neq NONE.
- On exit:* if **Interval Method** \neq NONE, BL(i, l) contains the lower limit of an $(100 \times \alpha)\%$ confidence interval for $B(i, l)$, for $i = 1, 2, \dots, IP$ and $l = 1, 2, \dots, NTAU$.
- If **Interval Method** = NONE, BL is not referenced.
- The method used for calculating the interval is controlled by the optional parameters **Interval Method** and **Bootstrap Interval Method**. The size of the interval, α , is controlled by the optional parameter **Significance Level**.
- 17: BU(IP,*) – REAL (KIND=nag_wp) array Output
- Note:** the second dimension of the array BU must be at least NTAU if **Interval Method** \neq NONE.
- On exit:* if **Interval Method** \neq NONE, BU(i, l) contains the upper limit of an $(100 \times \alpha)\%$ confidence interval for $B(i, l)$, for $i = 1, 2, \dots, IP$ and $l = 1, 2, \dots, NTAU$.
- If **Interval Method** = NONE, BU is not referenced.
- The method used for calculating the interval is controlled by the optional parameters **Interval Method** and **Bootstrap Interval Method**. The size of the interval, α is controlled by the optional parameter **Significance Level**.
- 18: CH(IP, IP,*) – REAL (KIND=nag_wp) array Output
- Note:** the last dimension of the array CH must be at least NTAU if **Interval Method** \neq NONE and **Matrix Returned** = COVARIANCE and at least NTAU + 1 if **Interval Method** \neq NONE, IID or BOOTSTRAP XY and **Matrix Returned** = H INVERSE.
- On exit:* depending on the supplied optional parameters, CH will either not be referenced, hold an estimate of the upper triangular part of the covariance matrix, Σ , or an estimate of the upper triangular parts of nJ_n and $n^{-1}H_n^{-1}$.
- If **Interval Method** = NONE or **Matrix Returned** = NONE, CH is not referenced.
- If **Interval Method** = BOOTSTRAP XY or IID and **Matrix Returned** = H INVERSE, CH is not referenced.
- Otherwise, for $i, j = 1, 2, \dots, IP, j \geq i$ and $l = 1, 2, \dots, NTAU$:
- If **Matrix Returned** = COVARIANCE, CH(i, j, l) holds an estimate of the covariance between $B(i, l)$ and $B(j, l)$.
- If **Matrix Returned** = H INVERSE, CH($i, j, 1$) holds an estimate of the (i, j) th element of nJ_n and CH($i, j, l+1$) holds an estimate of the (i, j) th element of $n^{-1}H_n^{-1}$, for $\tau = \text{TAU}(l)$.
- The method used for calculating Σ and H_n^{-1} is controlled by the optional parameter **Interval Method**.
- 19: RES(N,*) – REAL (KIND=nag_wp) array Output
- Note:** the second dimension of the array RES must be at least NTAU if **Return Residuals** = YES.
- On exit:* if **Return Residuals** = YES, RES(i, l) holds the (weighted) residuals, r_i , for $\tau = \text{TAU}(l)$, for $i = 1, 2, \dots, N$ and $l = 1, 2, \dots, NTAU$.
- If WEIGHT = 'W' and **Drop Zero Weights** = YES, the value of RES will be set to zero for observations with zero weights.
- If **Return Residuals** = NO, RES is not referenced.

- 20: IOPTS(*) – INTEGER array *Communication Array*
Note: the dimension of this array is dictated by the requirements of associated functions that must have been previously called. This array **must** be the same array passed as argument IOPTS in the previous call to G02ZKF.
On entry: optional parameter array, as initialized by a call to G02ZKF.
- 21: OPTS(*) – REAL (KIND=nag_wp) array *Communication Array*
Note: the dimension of this array is dictated by the requirements of associated functions that must have been previously called. This array **must** be the same array passed as argument OPTS in the previous call to G02ZKF.
On entry: optional parameter array, as initialized by a call to G02ZKF.
- 22: STATE(*) – INTEGER array *Communication Array*
Note: the actual argument supplied **must** be the array STATE supplied to the initialization routines G05KFF or G05KGF.
 The actual argument supplied **must** be the array STATE supplied to the initialization routines G05KFF or G05KGF.
 If **Interval Method** = BOOTSTRAP XY, STATE contains information about the selected random number generator. Otherwise STATE is not referenced.
- 23: INFO(NTAU) – INTEGER array *Output*
On exit: INFO(*i*) holds additional information concerning the model fitting and confidence limit calculations when $\tau = \text{TAU}(i)$.

Code Warning

- 0 Model fitted and confidence limits (if requested) calculated successfully
- 1 The routine did not converge. The returned values are based on the estimate at the last iteration. Try increasing **Iteration Limit** whilst calculating the parameter estimates or relaxing the definition of convergence by increasing **Tolerance**.
- 2 A singular matrix was encountered during the optimization. The model was not fitted for this value of τ .
- 4 Some truncation occurred whilst calculating the confidence limits for this value of τ . See Section 11 for details. The returned upper and lower limits may be narrower than specified.
- 8 The routine did not converge whilst calculating the confidence limits. The returned limits are based on the estimate at the last iteration. Try increasing **Iteration Limit**.
- 16 Confidence limits for this value of τ could not be calculated. The returned upper and lower limits are set to a large positive and large negative value respectively as defined by the optional parameter **Big**.

It is possible for multiple warnings to be applicable to a single model. In these cases the value returned in INFO is the sum of the corresponding individual nonzero warning codes.

- 24: IFAIL – INTEGER *Input/Output*
On entry: IFAIL must be set to 0, -1 or 1. If you are unfamiliar with this argument you should refer to Section 3.4 in How to Use the NAG Library and its Documentation for details.
 For environments where it might be inappropriate to halt program execution when an error is detected, the value -1 or 1 is recommended. If the output of error messages is undesirable, then the value 1 is recommended. Otherwise, if you are not familiar with this argument, the recommended value is 0. **When the value -1 or 1 is used it is essential to test the value of IFAIL on exit.**
On exit: IFAIL = 0 unless the routine detects an error or a warning has been flagged (see Section 6).

6 Error Indicators and Warnings

If on entry $IFAIL = 0$ or -1 , explanatory error messages are output on the current error message unit (as defined by X04AAF).

Errors or warnings detected by the routine:

$IFAIL = 11$

On entry, $SORDER = \langle value \rangle$.
Constraint: $SORDER = 1$ or 2 .

$IFAIL = 21$

On entry, $INTCPT = \langle value \rangle$ was an illegal value.

$IFAIL = 31$

On entry, $WEIGHT$ had an illegal value.

$IFAIL = 41$

On entry, $N = \langle value \rangle$.
Constraint: $N \geq 2$.

$IFAIL = 51$

On entry, $M = \langle value \rangle$.
Constraint: $M \geq 0$.

$IFAIL = 71$

On entry, $LDDAT = \langle value \rangle$ and $N = \langle value \rangle$.
Constraint: $LDDAT \geq N$.

$IFAIL = 72$

On entry, $LDDAT = \langle value \rangle$ and $M = \langle value \rangle$.
Constraint: $LDDAT \geq M$.

$IFAIL = 81$

On entry, $ISX(\langle value \rangle) = \langle value \rangle$.
Constraint: $ISX(i) = 0$ or 1 for all i .

$IFAIL = 91$

On entry, $IP = \langle value \rangle$ and $N = \langle value \rangle$.
Constraint: $1 \leq IP < N$.

$IFAIL = 92$

On entry, IP is not consistent with ISX or $INTCPT$: $IP = \langle value \rangle$, expected value = $\langle value \rangle$.

$IFAIL = 111$

On entry, $WT(\langle value \rangle) = \langle value \rangle$.
Constraint: $WT(i) \geq 0.0$ for all i .

$IFAIL = 112$

On entry, effective number of observations = $\langle value \rangle$.
Constraint: effective number of observations $\geq \langle value \rangle$.

IFAIL = 121

On entry, NTAU = $\langle value \rangle$.
Constraint: NTAU \geq 1.

IFAIL = 131

On entry, TAU($\langle value \rangle$) = $\langle value \rangle$ is invalid.

IFAIL = 201

On entry, either the option arrays have not been initialized or they have been corrupted.

IFAIL = 221

On entry, STATE vector has been corrupted or not initialized.

IFAIL = 231

A potential problem occurred whilst fitting the model(s).
Additional information has been returned in INFO.

IFAIL = -99

An unexpected error has been triggered by this routine. Please contact NAG.
See Section 3.9 in How to Use the NAG Library and its Documentation for further information.

IFAIL = -399

Your licence key may have expired or may not have been installed correctly.
See Section 3.8 in How to Use the NAG Library and its Documentation for further information.

IFAIL = -999

Dynamic memory allocation failed.
See Section 3.7 in How to Use the NAG Library and its Documentation for further information.

7 Accuracy

Not applicable.

8 Parallelism and Performance

G02QGF is threaded by NAG for parallel execution in multithreaded implementations of the NAG Library.

G02QGF makes calls to BLAS and/or LAPACK routines, which may be threaded within the vendor library used by this implementation. Consult the documentation for the vendor library for further information.

Please consult the X06 Chapter Introduction for information on how to control and interrogate the OpenMP environment used within this routine. Please also consult the Users' Note for your implementation for any additional implementation-specific information.

9 Further Comments

G02QGF allocates internally approximately the following elements of real storage: $13n + np + 3p^2 + 6p + 3(p + 1) \times \text{NTAU}$. If **Interval Method** = BOOTSTRAP XY then a further np elements are required, and this increases by $p \times \text{NTAU} \times \text{Bootstrap Iterations}$ if **Bootstrap Interval Method** = QUANTILE. Where possible, any user-supplied output arrays are used as workspace and so the amount actually allocated may be less. If SORDER = 2, WEIGHT = 'U',

INTCPT = 'N' and IP = M an internal copy of the input data is avoided and the amount of locally allocated memory is reduced by np .

10 Example

A quantile regression model is fitted to Engels 1857 study of household expenditure on food. The model regresses the dependent variable, household food expenditure, against two explanatory variables, a column of ones and household income. The model is fit for five different values of τ and the covariance matrix is estimated assuming Normal IID errors. Both the covariance matrix and the residuals are returned.

10.1 Program Text

```

Program g02qgfe

!      G02QGF Example Program Text

!      Mark 26 Release. NAG Copyright 2016.

!      .. Use Statements ..
      Use nag_library, Only: g02qgf, g02zkf, g02zlf, g05kff, nag_wp
!      .. Implicit None Statement ..
      Implicit None
!      .. Parameters ..
      Integer, Parameter          :: lseed = 1, nin = 5, nout = 6
!      .. Local Scalars ..
      Real (Kind=nag_wp)          :: df, rvalue
      Integer                     :: genid, i, ifail, ip, ivalue, j, l, &
                                   ldbl, lddat, ldres, liopts, lopts, &
                                   lstate, lwt, m, n, ntau, optype, &
                                   sorder, subid, tdch
      Character (1)               :: cl, weight
      Character (30)              :: cvalue, semeth
      Character (100)             :: optstr
!      .. Local Arrays ..
      Real (Kind=nag_wp), Allocatable :: b(:,,:), bl(:,,:), bu(:,,:), ch(:,,:), &
                                   dat(:,,:), opts(:), res(:,,:), tau(:), &
                                   wt(:), y(:)
      Integer, Allocatable         :: info(:), iopts(:), isx(:), state(:)
      Integer                     :: seed(lseed)
!      .. Intrinsic Procedures ..
      Intrinsic                   :: count, len_trim, min
!      .. Executable Statements ..
      Write (nout,*) 'G02QGF Example Program Results'
      Write (nout,*)
      Flush (nout)

!      Skip heading in data file
      Read (nin,*)

!      Read in the problem size
      Read (nin,*) sorder, cl, weight, n, m, ntau

!      Read in the data
      If (weight=='W' .Or. weight=='w') Then
         lwt = n
      Else
         lwt = 0
      End If
      Allocate (wt(lwt),isx(m),y(n),tau(ntau))

      If (sorder==1) Then
!      DAT(N,M)
         lddat = n
         Allocate (dat(lddat,m))
         If (lwt==0) Then
            Read (nin,*)(dat(i,1:m),y(i),i=1,n)
         Else

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        Read (nin,*)(dat(i,1:m),y(i),wt(i),i=1,n)
    End If
Else
!   DAT(M,N)
    lddat = m
    Allocate (dat(lddat,n))
    If (lwt==0) Then
        Read (nin,*)(dat(1:m,i),y(i),i=1,n)
    Else
        Read (nin,*)(dat(1:m,i),y(i),wt(i),i=1,n)
    End If
End If

!   Read in variable inclusion flags
Read (nin,*) isx(1:m)

!   Calculate IP
ip = count(isx(1:m)==1)
If (c1=='Y' .Or. c1=='y') Then
    ip = ip + 1
End If

!   Read in the quantiles required
Read (nin,*) tau(1:ntau)

    liopts = 100
    lopts = 100
    Allocate (iopts(liopts),opts(lopts))

!   Initialize the optional argument array
ifail = 0
Call g02zkf('INITIALIZE = G02QGF',iopts,liopts,opts,lopts,ifail)

c_lp: Do
!   Read in any optional arguments. Reads in to the end of
!   the input data, or until a blank line is reached
    ifail = 1
    Read (nin,99994,Iostat=ifail) optstr
    If (ifail/=0) Then
        Exit c_lp
    Else If (len_trim(optstr)==0) Then
        Exit c_lp
    End If

!   Set the supplied option
    ifail = 0
    Call g02zkf(optstr,iopts,liopts,opts,lopts,ifail)
End Do c_lp

!   Assume that no intervals or output matrices are required
!   unless the optional argument state differently
    ldbl = 0
    tdch = 0
    ldres = 0
    lstate = 0

!   Query the optional arguments to see what output is required
    ifail = 0
    Call g02zlf('INTERVAL METHOD',ivalue,rvalue,cvalue,optype,iopts,opts,
        & ifail)
    semeth = cvalue
    If (semeth/='NONE') Then
!   Require the intervals to be output
        ldbl = ip

        If (semeth=='BOOTSTRAP XY') Then
!   Need to find the length of the state array for the random
!   number generator

!   Read in the generator ID and a seed
        Read (nin,*) genid, subid, seed(1)

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!      Query the length of the state array
      Allocate (state(lstate))
      ifail = 0
      Call g05kff(genid,subid,seed,lseed,state,lstate,ifail)

!      Deallocate STATE so that it can reallocated later
      Deallocate (state)
End If

      ifail = 0
      Call g02zlf('MATRIX RETURNED',ivalue,rvalue,cvalue,optype,iopts,opts, &
        ifail)
      If (cvalue=='COVARIANCE') Then
        tdch = ntau
      Else If (cvalue=='H INVERSE') Then
        If (semeth=='BOOTSTRAP XY' .Or. semeth=='IID') Then
!          NB: If we are using bootstrap or IID errors then any request for
!          H INVERSE is ignored
          tdch = 0
        Else
          tdch = ntau + 1
        End If
      End If

      ifail = 0
      Call g02zlf('RETURN RESIDUALS',ivalue,rvalue,cvalue,optype,iopts,opts, &
        ifail)
      If (cvalue=='YES') Then
        ldres = n
      End If
End If

!      Allocate memory for output arrays
      Allocate (b(ip,ntau),info(ntau),bl(ldbl,ntau),bu(ldbl,ntau), &
        ch(ldbl,ldbl,tdch),state(lstate),res(ldres,ntau))

      If (lstate>0) Then
!        Doing bootstrap, so initialize the RNG
        ifail = 0
        Call g05kff(genid,subid,seed,lseed,state,lstate,ifail)
      End If

!      Call the model fitting routine
      ifail = -1
      Call g02qgf(sorder,c1,weight,n,m,dat,lddat,isx,ip,y,wt,ntau,tau,df,b,bl, &
        bu,ch,res,iopts,opts,state,info,ifail)
      If (ifail/=0) Then
        If (ifail==231) Then
          Write (nout,*) 'Additional error information (INFO): ', info(1:ntau)
        Else
          Go To 100
        End If
      End If

!      Display the parameter estimates
      Do l = 1, ntau
        Write (nout,99999) 'Quantile: ', tau(l)
        Write (nout,*)
        If (ldbl>0) Then
          Write (nout,*) '          Lower    Parameter    Upper'
          Write (nout,*) '          Limit      Estimate      Limit'
        Else
          Write (nout,*) '          Parameter'
          Write (nout,*) '          Estimate'
        End If
        Do j = 1, ip
          If (ldbl>0) Then
            Write (nout,99998) j, bl(j,l), b(j,l), bu(j,l)
          Else
            Write (nout,99998) j, b(j,l)
          End If
        End Do
      End Do

```

```

      End If
    End Do
    Write (nout,*)
    If (tdch==ntau) Then
      Write (nout,*) 'Covariance matrix'
      Do i = 1, ip
        Write (nout,99997) ch(1:i,i,1)
      End Do
      Write (nout,*)
    Else If (tdch==ntau+1) Then
      Write (nout,*) 'J'
      Do i = 1, ip
        Write (nout,99997) ch(1:i,i,1)
      End Do
      Write (nout,*)
      Write (nout,*) 'H inverse'
      Do i = 1, ip
        Write (nout,99997) ch(1:i,i,1+1)
      End Do
      Write (nout,*)
    End If
    Write (nout,*)
  End Do

  If (ldres>0) Then
    Write (nout,*) 'First 10 Residuals'
    Write (nout,*) '
    Write (nout,99995) 'Obs.', tau(1:ntau)
    Do i = 1, min(n,10)
      Write (nout,99996) i, res(i,1:ntau)
    End Do
  Else
    Write (nout,*) 'Residuals not returned'
  End If
  Write (nout,*)

100  Continue

99999 Format (1X,A,F6.3)
99998 Format (1X,I3,3(3X,F7.3))
99997 Format (1X,10(E10.3,3X))
99996 Format (2X,I3,10(1X,F10.5))
99995 Format (1X,A,10(3X,F6.3,2X))
99994 Format (A100)
      End Program g02qgfe

```

10.2 Program Data

G02QGF Example Program Data

1 'Y' 'U'	235	1	5	:: SORDER,C1,WEIGHT,N,M,NTAU
420.1577	255.8394	800.7990	572.0807	643.3571 459.8177
541.4117	310.9587	1245.6964	907.3969	2551.6615 863.9199
901.1575	485.6800	1201.0002	811.5776	1795.3226 831.4407
639.0802	402.9974	634.4002	427.7975	1165.7734 534.7610
750.8756	495.5608	956.2315	649.9985	815.6212 392.0502
945.7989	633.7978	1148.6010	860.6002	1264.2066 934.9752
829.3979	630.7566	1768.8236	1143.4211	1095.4056 813.3081
979.1648	700.4409	2822.5330	2032.6792	447.4479 263.7100
1309.8789	830.9586	922.3548	590.6183	1178.9742 769.0838
1492.3987	815.3602	2293.1920	1570.3911	975.8023 630.5863
502.8390	338.0014	627.4726	483.4800	1017.8522 645.9874
616.7168	412.3613	889.9809	600.4804	423.8798 319.5584
790.9225	520.0006	1162.2000	696.2021	558.7767 348.4518
555.8786	452.4015	1197.0794	774.7962	943.2487 614.5068
713.4412	512.7201	530.7972	390.5984	1348.3002 662.0096
838.7561	658.8395	1142.1526	612.5619	2340.6174 1504.3708
535.0766	392.5995	1088.0039	708.7622	587.1792 406.2180
596.4408	443.5586	484.6612	296.9192	1540.9741 692.1689
924.5619	640.1164	1536.0201	1071.4627	1115.8481 588.1371
487.7583	333.8394	678.8974	496.5976	1044.6843 511.2609

```

692.6397 466.9583 671.8802 503.3974 1389.7929 700.5600
997.8770 543.3969 690.4683 357.6411 2497.7860 1301.1451
506.9995 317.7198 860.6948 430.3376 1585.3809 879.0660
654.1587 424.3209 873.3095 624.6990 1862.0438 912.8851
933.9193 518.9617 894.4598 582.5413 2008.8546 1509.7812
433.6813 338.0014 1148.6470 580.2215 697.3099 484.0605
587.5962 419.6412 926.8762 543.8807 571.2517 399.6703
896.4746 476.3200 839.0414 588.6372 598.3465 444.1001
454.4782 386.3602 829.4974 627.9999 461.0977 248.8101
584.9989 423.2783 1264.0043 712.1012 977.1107 527.8014
800.7990 503.3572 1937.9771 968.3949 883.9849 500.6313
502.4369 354.6389 698.8317 482.5816 718.3594 436.8107
713.5197 497.3182 920.4199 593.1694 543.8971 374.7990
906.0006 588.5195 1897.5711 1033.5658 1587.3480 726.3921
880.5969 654.5971 891.6824 693.6795 4957.8130 1827.2000
796.8289 550.7274 889.6784 693.6795 969.6838 523.4911
854.8791 528.3770 1221.4818 761.2791 419.9980 334.9998
1167.3716 640.4813 544.5991 361.3981 561.9990 473.2009
523.8000 401.3204 1031.4491 628.4522 689.5988 581.2029
670.7792 435.9990 1462.9497 771.4486 1398.5203 929.7540
377.0584 276.5606 830.4353 757.1187 820.8168 591.1974
851.5430 588.3488 975.0415 821.5970 875.1716 637.5483
1121.0937 664.1978 1337.9983 1022.3202 1392.4499 674.9509
625.5179 444.8602 867.6427 679.4407 1256.3174 776.7589
805.5377 462.8995 725.7459 538.7491 1362.8590 959.5170
558.5812 377.7792 989.0056 679.9981 1999.2552 1250.9643
884.4005 553.1504 1525.0005 977.0033 1209.4730 737.8201
1257.4989 810.8962 672.1960 561.2015 1125.0356 810.6772
2051.1789 1067.9541 923.3977 728.3997 1827.4010 983.0009
1466.3330 1049.8788 472.3215 372.3186 1014.1540 708.8968
730.0989 522.7012 590.7601 361.5210 880.3944 633.1200
2432.3910 1424.8047 831.7983 620.8006 873.7375 631.7982
940.9218 517.9196 1139.4945 819.9964 951.4432 608.6419
1177.8547 830.9586 507.5169 360.8780 473.0022 300.9999
1222.5939 925.5795 576.1972 395.7608 601.0030 377.9984
1519.5811 1162.0024 696.5991 442.0001 713.9979 397.0015
687.6638 383.4580 650.8180 404.0384 829.2984 588.5195
953.1192 621.1173 949.5802 670.7993 959.7953 681.7616
953.1192 621.1173 497.1193 297.5702 1212.9613 807.3603
953.1192 621.1173 570.1674 353.4882 958.8743 696.8011
939.0418 548.6002 724.7306 383.9376 1129.4431 811.1962
1283.4025 745.2353 408.3399 284.8008 1943.0419 1305.7201
1511.5789 837.8005 638.6713 431.1000 539.6388 442.0001
1342.5821 795.3402 1225.7890 801.3518 463.5990 353.6013
511.7980 418.5976 715.3701 448.4513 562.6400 468.0008
689.7988 508.7974 800.4708 577.9111 736.7584 526.7573
1532.3074 883.2780 975.5974 570.5210 1415.4461 890.2390
1056.0808 742.5276 1613.7565 865.3205 2208.7897 1318.8033
387.3195 242.3202 608.5019 444.5578 636.0009 331.0005
387.3195 242.3202 958.6634 680.4198 759.4010 416.4015
410.9987 266.0010 835.9426 576.2779 1078.8382 596.8406
499.7510 408.4992 1024.8177 708.4787 748.6413 429.0399
832.7554 614.7588 1006.4353 734.2356 987.6417 619.6408
614.9986 385.3184 726.0000 433.0010 788.0961 400.7990
887.4658 515.6200 494.4174 327.4188 1020.0225 775.0209
1595.1611 1138.1620 776.5958 485.5198 1230.9235 772.7611
1807.9520 993.9630 415.4407 305.4390 440.5174 306.5191
541.2006 299.1993 581.3599 468.0008 743.0772 522.6019
1057.6767 750.3202 :: End of X,Y (in three set of columns)
1 :: ISX
0.10 0.25 0.50 0.75 0.90 :: TAU
Return Residuals = Yes
Matrix Returned = Covariance
Interval Method = IID

```

10.3 Program Results

G02QGF Example Program Results

Quantile: 0.100

	Lower Limit	Parameter Estimate	Upper Limit
1	74.946	110.142	145.337
2	0.370	0.402	0.433

Covariance matrix

0.319E+03
-0.254E+00 0.259E-03

Quantile: 0.250

	Lower Limit	Parameter Estimate	Upper Limit
1	64.232	95.483	126.735
2	0.446	0.474	0.502

Covariance matrix

0.252E+03
-0.200E+00 0.204E-03

Quantile: 0.500

	Lower Limit	Parameter Estimate	Upper Limit
1	55.399	81.482	107.566
2	0.537	0.560	0.584

Covariance matrix

0.175E+03
-0.140E+00 0.142E-03

Quantile: 0.750

	Lower Limit	Parameter Estimate	Upper Limit
1	41.372	62.396	83.421
2	0.625	0.644	0.663

Covariance matrix

0.114E+03
-0.907E-01 0.923E-04

Quantile: 0.900

	Lower Limit	Parameter Estimate	Upper Limit
1	26.829	67.351	107.873
2	0.650	0.686	0.723

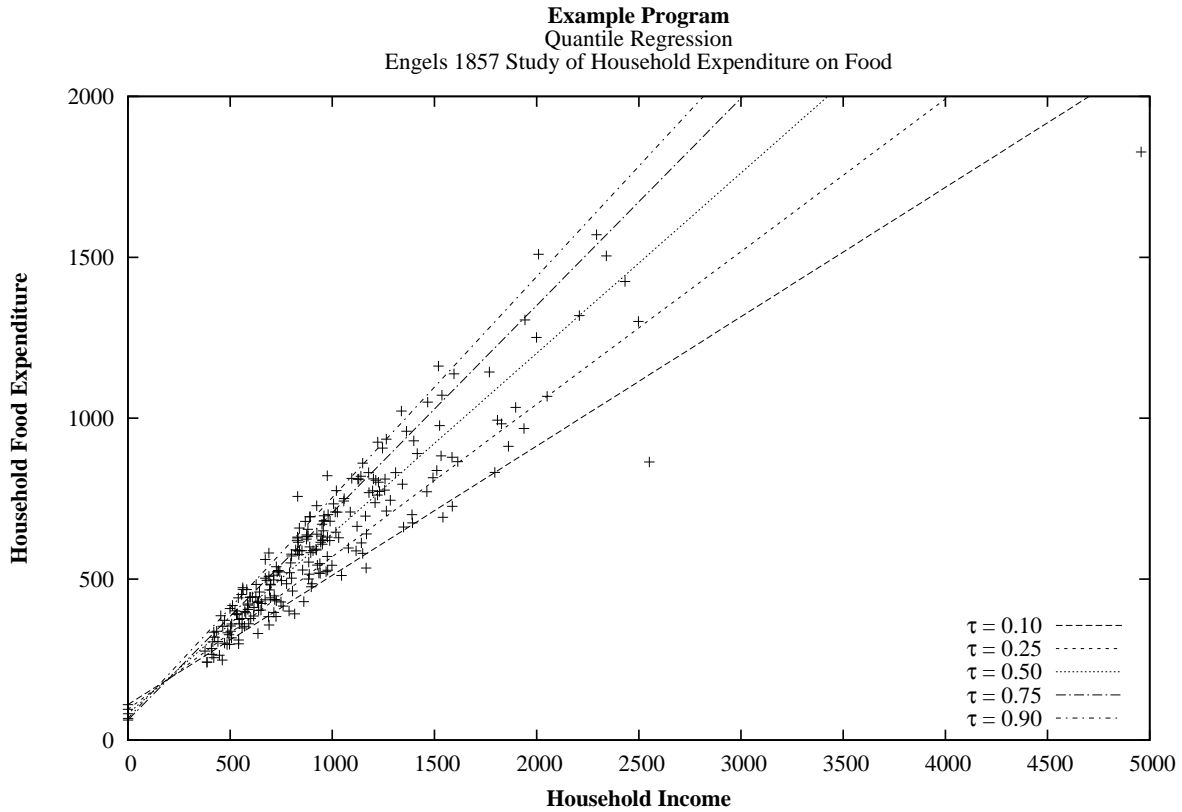
Covariance matrix

0.423E+03
-0.337E+00 0.343E-03

First 10 Residuals

	0.100	0.250	0.500	0.750	0.900
Obs.					
1	-23.10718	-38.84219	-61.00711	-77.14462	-99.86551
2	140.20549	96.93582	42.00636	-6.04177	-44.85812
3	91.19725	59.31654	17.93924	-16.90993	-49.06884
4	-16.70358	-41.20981	-73.81193	-100.11463	-127.96277

5	296.77717	221.32470	128.09970	42.75414	-14.87476
6	-271.39185	-441.31464	-646.95350	-841.78309	-954.63488
7	13.48419	-37.04518	-100.61322	-157.07478	-200.13481
8	218.91527	146.69601	57.31834	-24.28017	-80.01908
9	0.00000	-115.21109	-255.74639	-387.16920	-468.03911
10	36.09526	4.52393	-36.48522	-70.97584	-102.95390



11 Algorithmic Details

By the addition of slack variables the minimization (1) can be reformulated into the linear programming problem

$$\underset{(u,v,\beta) \in \mathbb{R}_+^n \times \mathbb{R}_+^n \times \mathbb{R}^p}{\text{minimize}} \quad \tau e^T u + (1 - \tau) e^T v \quad \text{subject to} \quad y = X\beta + u - v \quad (2)$$

and its associated dual

$$\underset{d}{\text{maximize}} \quad y^T d \quad \text{subject to} \quad X^T d = 0, d \in [\tau - 1, \tau]^n \quad (3)$$

where e is a vector of n 1s. Setting $a = d + (1 - \tau)e$ gives the equivalent formulation

$$\underset{a}{\text{maximize}} \quad y^T a \quad \text{subject to} \quad X^T a = (1 - \tau) X^T e, a \in [0, 1]^n. \quad (4)$$

The algorithm introduced by Portnoy and Koenker (1997) and used by G02QGF, uses the primal-dual formulation expressed in equations (2) and (4) along with a logarithmic barrier function to obtain estimates for β . The algorithm is based on the predictor-corrector algorithm of Mehrotra (1992) and further details can be obtained from Portnoy and Koenker (1997) and Koenker (2005). A good description of linear programming, interior point algorithms, barrier functions and Mehrotra's predictor-corrector algorithm can be found in Nocedal and Wright (1999).

11.1 Interior Point Algorithm

In this section a brief description of the interior point algorithm used to estimate the model parameters is presented. It should be noted that there are some differences in the equations given here – particularly (7) and (9) – compared to those given in Koenker (2005) and Portnoy and Koenker (1997).

11.1.1 Central path

Rather than optimize (4) directly, an additional slack variable s is added and the constraint $a \in [0, 1]^n$ is replaced with $a + s = e, a_i \geq 0, s_i \geq 0$, for $i = 1, 2, \dots, n$.

The positivity constraint on a and s is handled using the logarithmic barrier function

$$B(a, s, \mu) = y^T a + \mu \sum_{i=1}^n (\log a_i + \log s_i).$$

The primal-dual form of the problem is used giving the Lagrangian

$$L(a, s, \beta, u, \mu) = B(a, s, \mu) - \beta^T (X^T a - (1 - \tau)X^T e) - u^T (a + s - e)$$

whose central path is described by the following first order conditions

$$\begin{aligned} X^T a &= (1 - \tau)X^T e \\ a + s &= e \\ X\beta + u - v &= y \\ S U e &= \mu e \\ A V e &= \mu e \end{aligned} \tag{5}$$

where A denotes the diagonal matrix with diagonal elements given by a , similarly with S, U and V . By enforcing the inequalities on s and a strictly, i.e., $a_i > 0$ and $s_i > 0$ for all i we ensure that A and S are positive definite diagonal matrices and hence A^{-1} and S^{-1} exist.

Rather than applying Newton's method to the system of equations given in (5) to obtain the step directions $\delta_\beta, \delta_a, \delta_s, \delta_u$ and δ_v , Mehrotra substituted the steps directly into (5) giving the augmented system of equations

$$\begin{aligned} X^T(a + \delta_a) &= (1 - \tau)X^T e \\ (a + \delta_a) + (s + \delta_s) &= e \\ X(\beta + \delta_\beta) + (u + \delta_u) - (v + \delta_v) &= y \\ (S + \Delta_s)(U + \Delta_u)e &= \mu e \\ (A + \Delta_a)(V + \Delta_v)e &= \mu e \end{aligned} \tag{6}$$

where $\Delta_a, \Delta_s, \Delta_u$ and Δ_v denote the diagonal matrices with diagonal elements given by $\delta_a, \delta_s, \delta_u$ and δ_v respectively.

11.1.2 Affine scaling step

The affine scaling step is constructed by setting $\mu = 0$ in (5) and applying Newton's method to obtain an intermediate set of step directions

$$\begin{aligned} (X^T W X) \delta_\beta &= X^T W (y - X\beta) + (\tau - 1)X^T e + X^T a \\ \delta_a &= W(y - X\beta - X\delta_\beta) \\ \delta_s &= -\delta_a \\ \delta_u &= S^{-1}U\delta_a - Ue \\ \delta_v &= A^{-1}V\delta_s - Ve \end{aligned} \tag{7}$$

where $W = (S^{-1}U + A^{-1}V)^{-1}$.

Initial step sizes for the primal ($\hat{\gamma}_P$) and dual ($\hat{\gamma}_D$) parameters are constructed as

$$\begin{aligned} \hat{\gamma}_P &= \sigma \min \left\{ \min_{i, \delta_{a_i} < 0} \{a_i / \delta_{a_i}\}, \min_{i, \delta_{s_i} < 0} \{s_i / \delta_{s_i}\} \right\} \\ \hat{\gamma}_D &= \sigma \min \left\{ \min_{i, \delta_{u_i} < 0} \{u_i / \delta_{u_i}\}, \min_{i, \delta_{v_i} < 0} \{v_i / \delta_{v_i}\} \right\} \end{aligned} \tag{8}$$

where σ is a user-supplied scaling factor. If $\hat{\gamma}_P \times \hat{\gamma}_D \geq 1$ then the nonlinearity adjustment, described in Section 11.1.3, is not made and the model parameters are updated using the current step size and directions.

11.1.3 Nonlinearity Adjustment

In the nonlinearity adjustment step a new estimate of μ is obtained by letting

$$\hat{g}(\hat{\gamma}_P, \hat{\gamma}_D) = (s + \hat{\gamma}_P \delta_s)^T (u + \hat{\gamma}_D \delta_u) + (a + \hat{\gamma}_P \delta_a)^T (v + \hat{\gamma}_D \delta_v)$$

and estimating μ as

$$\mu = \left(\frac{\hat{g}(\hat{\gamma}_P, \hat{\gamma}_D)}{\hat{g}(0, 0)} \right)^3 \frac{\hat{g}(0, 0)}{2n}.$$

This estimate, along with the nonlinear terms (Δu , Δs , Δa and Δv) from (6) are calculated using the values of δ_a , δ_s , δ_u and δ_v obtained from the affine scaling step.

Given an updated estimate for μ and the nonlinear terms the system of equations

$$\begin{aligned} (X^T W X) \delta_\beta &= X^T W (y - X\beta + \mu(S^{-1} - A^{-1})e + S^{-1} \Delta_s \Delta_u e - A^{-1} \Delta_a \Delta_v e) + (\tau - 1)X^T e + X^T a \\ \delta_a &= W(y - X\beta - X\delta_\beta + \mu(S^{-1} - A^{-1})e) \\ \delta_s &= -\delta_a \\ \delta_u &= \mu S^{-1} e + S^{-1} U \delta_a - Ue - S^{-1} \Delta_s \Delta_u e \\ \delta_v &= \mu A^{-1} e + A^{-1} V \delta_s - Ve - A^{-1} \Delta_a \Delta_v e \end{aligned} \tag{9}$$

are solved and updated values for δ_β , δ_a , δ_s , δ_u , δ_v , $\hat{\gamma}_P$ and $\hat{\gamma}_D$ calculated.

11.1.4 Update and convergence

At each iteration the model parameters (β, a, s, u, v) are updated using step directions, $(\delta_\beta, \delta_a, \delta_s, \delta_u, \delta_v)$ and step lengths $(\hat{\gamma}_P, \hat{\gamma}_D)$.

Convergence is assessed using the duality gap, that is, the differences between the objective function in the primal and dual formulations. For any feasible point (u, v, s, a) the duality gap can be calculated from equations (2) and (3) as

$$\begin{aligned} \tau e^T u + (1 - \tau)e^T v - d^T y &= \tau e^T u + (1 - \tau)e^T v - (a - (1 - \tau)e)^T y \\ &= s^T u + a^T v \\ &= e^T u - a^T y + (1 - \tau)e^T X\beta \end{aligned}$$

and the optimization terminates if the duality gap is smaller than the tolerance supplied in the optional parameter **Tolerance**.

11.1.5 Additional information

Initial values are required for the parameters a, s, u, v and β . If not supplied by the user, initial values for β are calculated from a least squares regression of y on X . This regression is carried out by first constructing the cross-product matrix $X^T X$ and then using a pivoted QR decomposition as performed by F08BFF (DGEQP3). In addition, if the cross-product matrix is not of full rank, a rank reduction is carried out and, rather than using the full design matrix, X , a matrix formed from the first p -rank columns of XP is used instead, where P is the pivot matrix used during the QR decomposition. Parameter estimates, confidence intervals and the rows and columns of the matrices returned in the argument CH (if any) are set to zero for variables dropped during the rank-reduction. The rank reduction step is performed irrespective of whether initial values are supplied by the user.

Once initial values have been obtained for β , the initial values for u and v are calculated from the residuals. If $|r_i| < \epsilon_u$ then a value of $\pm \epsilon_u$ is used instead, where ϵ_u is supplied in the optional parameter **Epsilon**. The initial values for the a and s are always set to $1 - \tau$ and τ respectively.

The solution for δ_β in both (7) and (9) is obtained using a Bunch–Kaufman decomposition, as implemented in F07MDF (DSYTRF).

11.2 Calculation of Covariance Matrix

G02QGF supplies four methods to calculate the covariance matrices associated with the parameter estimates for β . This section gives some additional detail on three of the algorithms, the fourth, (which uses bootstrapping), is described in Section 3.

(i) Independent, identically distributed (IID) errors

When assuming IID errors, the covariance matrices depend on the sparsity, $s(\tau)$, which G02QGF estimates as follows:

- (a) Let r_i denote the residuals from the original quantile regression, that is $r_i = y_i - x_i^T \hat{\beta}$.
- (b) Drop any residual where $|r_i|$ is less than ϵ_u , supplied in the optional parameter **Epsilon**.
- (c) Sort and relabel the remaining residuals in ascending order, by absolute value, so that $\epsilon_u < |r_1| < |r_2| < \dots$.
- (d) Select the first l values where $l = h_n n$, for some bandwidth h_n .
- (e) Sort and relabel these l residuals again, so that $r_1 < r_2 < \dots < r_l$ and regress them against a design matrix with two columns ($p = 2$) and rows given by $x_i = \{1, i/(n - p)\}$ using quantile regression with $\tau = 0.5$.
- (f) Use the resulting estimate of the slope as an estimate of the sparsity.

(ii) Powell Sandwich

When using the Powell Sandwich to estimate the matrix H_n , the quantity

$$c_n = \min(\sigma_r, (q_{r3} - q_{r1})/1.34) \times (\Phi^{-1}(\tau + h_n) - \Phi^{-1}(\tau - h_n))$$

is calculated. Dependent on the value of τ and the method used to calculate the bandwidth (h_n), it is possible for the quantities $\tau \pm h_n$ to be too large or small, compared to *machine precision* (ϵ). More specifically, when $\tau - h_n \leq \sqrt{\epsilon}$, or $\tau + h_n \geq 1 - \sqrt{\epsilon}$, a warning flag is raised in INFO, the value is truncated to $\sqrt{\epsilon}$ or $1 - \sqrt{\epsilon}$ respectively and the covariance matrix calculated as usual.

(iii) Hendricks–Koenker Sandwich

The Hendricks–Koenker Sandwich requires the calculation of the quantity $d_i = x_i^T (\hat{\beta}(\tau + h_n) - \hat{\beta}(\tau - h_n))$. As with the Powell Sandwich, in cases where $\tau - h_n \leq \sqrt{\epsilon}$, or $\tau + h_n \geq 1 - \sqrt{\epsilon}$, a warning flag is raised in INFO, the value truncated to $\sqrt{\epsilon}$ or $1 - \sqrt{\epsilon}$ respectively and the covariance matrix calculated as usual.

In addition, it is required that $d_i > 0$, in this method. Hence, instead of using $2h_n/d_i$ in the calculation of H_n , $\max(2h_n/(d_i + \epsilon_u), 0)$ is used instead, where ϵ_u is supplied in the optional parameter **Epsilon**.

12 Optional Parameters

Several optional parameters in G02QGF control aspects of the optimization algorithm, methodology used, logic or output. Their values are contained in the arrays IOPTS and OPTS; these must be initialized before calling G02QGF by first calling G02ZKF with OPTSTR set to **Initialize** = G02QGF.

Each optional parameter has an associated default value; to set any of them to a non-default value, use G02ZKF. The current value of an optional parameter can be queried using G02ZLF.

The remainder of this section can be skipped if you wish to use the default values for all optional parameters.

The following is a list of the optional parameters available. A full description of each optional parameter is provided in Section 12.1.

Band Width Alpha

Band Width Method

Big

Bootstrap Interval Method**Bootstrap Iterations****Bootstrap Monitoring****Calculate Initial Values****Defaults****Drop Zero Weights****Epsilon****Interval Method****Iteration Limit****Matrix Returned****Monitoring****QR Tolerance****Return Residuals****Sigma****Significance Level****Tolerance****Unit Number****12.1 Description of the Optional Parameters**

For each option, we give a summary line, a description of the optional parameter and details of constraints.

The summary line contains:

the keywords, where the minimum abbreviation of each keyword is underlined (if no characters of an optional qualifier are underlined, the qualifier may be omitted);

a parameter value, where the letters a , i and r denote options that take character, integer and real values respectively;

the default value, where the symbol ϵ is a generic notation for *machine precision* (see X02AJF).

Keywords and character values are case and white space insensitive.

Band Width Alpha r Default = 1.0

A multiplier used to construct the parameter α_b used when calculating the Sheather–Hall bandwidth (see Section 3), with $\alpha_b = (1 - \alpha) \times \mathbf{Band\ Width\ Alpha}$. Here, α is the **Significance Level**.

Constraint: **Band Width Alpha** > 0.0.

Band Width Method a Default = 'SHEATHER HALL'

The method used to calculate the bandwidth used in the calculation of the asymptotic covariance matrix Σ and H^{-1} if **Interval Method** = HKS, KERNEL or IID (see Section 3).

Constraint: **Band Width Method** = SHEATHER HALL or BOFINGER.

Big r Default = 10.0²⁰

This parameter should be set to something larger than the biggest value supplied in DAT and Y.

Constraint: **Big** > 0.0.

Bootstrap Interval Method*a*

Default = QUANTILE

If **Interval Method** = BOOTSTRAP XY, **Bootstrap Interval Method** controls how the confidence intervals are calculated from the bootstrap estimates.

Bootstrap Interval Method = T

t intervals are calculated. That is, the covariance matrix, $\Sigma = \{\sigma_{ij} : i, j = 1, 2, \dots, p\}$ is calculated from the bootstrap estimates and the limits calculated as $\beta_i \pm t_{(n-p, (1+\alpha)/2)} \sigma_{ii}$ where $t_{(n-p, (1+\alpha)/2)}$ is the $(1 + \alpha)/2$ percentage point from a Student's t distribution on $n - p$ degrees of freedom, n is the effective number of observations and α is given by the optional parameter **Significance Level**.

Bootstrap Interval Method = QUANTILE

Quantile intervals are calculated. That is, the upper and lower limits are taken as the $(1 + \alpha)/2$ and $(1 - \alpha)/2$ quantiles of the bootstrap estimates, as calculated using G01AMF.

Constraint: **Bootstrap Interval Method** = T or QUANTILE.

Bootstrap Iterations*i*

Default = 100

The number of bootstrap samples used to calculate the confidence limits and covariance matrix (if requested) when **Interval Method** = BOOTSTRAP XY.

Constraint: **Bootstrap Iterations** > 1.

Bootstrap Monitoring*a*

Default = NO

If **Bootstrap Monitoring** = YES and **Interval Method** = BOOTSTRAP XY, then the parameter estimates for each of the bootstrap samples are displayed. This information is sent to the unit number specified by **Unit Number**.

Constraint: **Bootstrap Monitoring** = YES or NO.

Calculate Initial Values*a*

Default = YES

If **Calculate Initial Values** = YES then the initial values for the regression parameters, β , are calculated from the data. Otherwise they must be supplied in B.

Constraint: **Calculate Initial Values** = YES or NO.

Defaults

This special keyword is used to reset all optional parameters to their default values.

Drop Zero Weights*a*

Default = YES

If a weighted regression is being performed and **Drop Zero Weights** = YES then observations with zero weight are dropped from the analysis. Otherwise such observations are included.

Constraint: **Drop Zero Weights** = YES or NO.

Epsilon*r*Default = $\sqrt{\epsilon}$

ϵ_u , the tolerance used when calculating the covariance matrix and the initial values for u and v . For additional details see Section 11.2 and Section 11.1.5 respectively.

Constraint: **Epsilon** ≥ 0.0 .

Interval Method*a*

Default = IID

The value of **Interval Method** controls whether confidence limits are returned in BL and BU and how these limits are calculated. This parameter also controls how the matrices returned in CH are calculated.

Interval Method = NONE

No limits are calculated and BL, BU and CH are not referenced.

Interval Method = KERNEL

The Powell Sandwich method with a Gaussian kernel is used.

Interval Method = HKS

The Hendricks–Koenker Sandwich is used.

Interval Method = IID

The errors are assumed to be identical, and independently distributed.

Interval Method = BOOTSTRAP XY

A bootstrap method is used, where sampling is done on the pair (y_i, x_i) . The number of bootstrap samples is controlled by the parameter **Bootstrap Iterations** and the type of interval constructed from the bootstrap samples is controlled by **Bootstrap Interval Method**.

Constraint: **Interval Method** = NONE, KERNEL, HKS, IID or BOOTSTRAP XY.

Iteration Limit i

Default = 100

The maximum number of iterations to be performed by the interior point optimization algorithm.

Constraint: **Iteration Limit** > 0.

Matrix Returned a

Default = NONE

The value of **Matrix Returned** controls the type of matrices returned in CH. If **Interval Method** = NONE, this parameter is ignored and CH is not referenced. Otherwise:

Matrix Returned = NONE

No matrices are returned and CH is not referenced.

Matrix Returned = COVARIANCE

The covariance matrices are returned.

Matrix Returned = H INVERSE

If **Interval Method** = KERNEL or HKS, the matrices J and H^{-1} are returned. Otherwise no matrices are returned and CH is not referenced.

The matrices returned are calculated as described in Section 3, with the algorithm used specified by **Interval Method**. In the case of **Interval Method** = BOOTSTRAP XY the covariance matrix is calculated directly from the bootstrap estimates.

Constraint: **Matrix Returned** = NONE, COVARIANCE or H INVERSE.

Monitoring a

Default = NO

If **Monitoring** = YES then the duality gap is displayed at each iteration of the interior point optimization algorithm. In addition, the final estimates for β are also displayed.

The monitoring information is sent to the unit number specified by **Unit Number**.

Constraint: **Monitoring** = YES or NO.

QR Tolerance r Default = $\epsilon^{0.9}$

The tolerance used to calculate the rank, k , of the $p \times p$ cross-product matrix, $X^T X$. Letting Q be the orthogonal matrix obtained from a QR decomposition of $X^T X$, then the rank is calculated by comparing Q_{ii} with $Q_{11} \times$ **QR Tolerance**.

If the cross-product matrix is rank deficient, then the parameter estimates for the $p - k$ columns with the smallest values of Q_{ii} are set to zero, along with the corresponding entries in BL, BU and CH, if returned. This is equivalent to dropping these variables from the model. Details on the QR decomposition used can be found in F08BFF (DGEQP3).

Constraint: **QR Tolerance** > 0.0.

Return Residuals a

Default = NO

If **Return Residuals** = YES, the residuals are returned in RES. Otherwise RES is not referenced.

Constraint: **Return Residuals** = YES or NO.

Sigma r

Default = 0.99995

The scaling factor used when calculating the affine scaling step size (see equation (8)).

Constraint: $0.0 < \mathbf{Sigma} < 1.0$.

Significance Level r

Default = 0.95

α , the size of the confidence interval whose limits are returned in BL and BU.

Constraint: $0.0 < \mathbf{Significance Level} < 1.0$.

Tolerance r Default = $\sqrt{\epsilon}$

Convergence tolerance. The optimization is deemed to have converged if the duality gap is less than **Tolerance** (see Section 11.1.4).

Constraint: **Tolerance** > 0.0 .

Unit Number i

Default taken from X04ABF

The unit number to which any monitoring information is sent.

Constraint: **Unit Number** > 1 .

13 Description of Monitoring Information

See the description of the optional argument **Monitoring**.
